

## A Willow tree construction

Now, we give details on the constructions of the willow tree for three popular stochastic processes, such as the geometric Brownian motion (GBM), Merton's jump-diffusion (MJD) (Merton, 1976) and Heston's stochastic volatility (HSV) (Heston, 1993) models.

### A.1 Geometric Brownian motion

Assume the underlying asset price  $S_t$  under the  $\mathbb{Q}$  measure are governed by the stochastic differential equation (SDE)

$$dS_t = rS_t dt + \sigma S_t dW_t \quad (\text{A.24})$$

where  $r$  is the risk-free interest rate,  $W_t$  is the standard Brownian motion,  $\sigma$  is the constant volatility. The value of  $S_t$  at  $t$  can be estimated as

$$S_t = S_0 \exp \left\{ \left( r - \frac{\sigma^2}{2} \right) t + \sigma W_t \right\}. \quad (\text{A.25})$$

Thus, the discrete asset price  $S_i^n$  at  $t_n$  can be estimated as

$$S_i^n = S_0 \exp \left\{ \left( r - \frac{\sigma^2}{2} \right) t_n + \sigma \sqrt{t_n} z_i \right\}, \quad \text{for } i = 1, \dots, m. \quad (\text{A.26})$$

where  $z_i$  is the discrete value chosen from the standard normal distribution following the sample strategy in Xu et al. (2013). According to Xu et al. (2013), a sequence of  $\{(z_i, \hat{q}_i)\}$ ,  $i = 1, 2, \dots, m$ , is generated to approximate the standard normal distribution, where  $z_i$  is some discrete value of the standard normal distribution and  $\hat{q}_i$  is the corresponding probability of  $z_i$ .

The transition probability from  $S_i^n$  to  $S_j^{n+1}$ ,  $p_{ij}^n$ , can then be estimated as (Lu and Xu, 2017)

$$p_{ij}^n = P(Y_j^{n+1} | Y_i^n) = \int_{c_j^{n+1}}^{c_{j+1}^{n+1}} f(y | Y_i^n) dy, \quad \text{for } i, j = 1, \dots, m,$$

where  $Y_i^n \equiv \sqrt{t_n} z_i$ ,  $c_j^{n+1} = (Y_j^{n+1} + Y_{j-1}^{n+1})/2$ ,  $c_{j+1}^{n+1} = (Y_{j+1}^{n+1} + Y_j^{n+1})/2$ ,  $c_1^{n+1} = -\infty$ ,  $c_{m+1}^{n+1} = +\infty$  for  $j = 1, 2, \dots, m$ , and  $f(y | Y_i^n)$  is the conditional probability density function for a normally distributed random variable  $y$  at  $t_{n+1}$ , given  $Y_i^n$ , i.e.,

$$f(y | Y_i^n) = \frac{1}{\sqrt{2\pi\Delta t}} \exp \left\{ -\frac{(y - Y_i^n)^2}{2\Delta t} \right\},$$

for  $n = 1, \dots, N - 1$ . The transition probability from  $S_0$  to  $S_j^1$ ,  $q_j$ , can be determined by

$$q_j = P(Y_j^1 | Y^0) = \int_{c_j^1}^{c_{j+1}^1} f(y) dy,$$

where  $f(y) = \frac{1}{\sqrt{2\pi\Delta t}} \exp \left\{ -\frac{y^2}{2\Delta t} \right\}$ .

## A.2 Merton's jump-diffusion model

Assume the underlying asset price  $S_t$  follows a jump-diffusion process (Merton, 1976) as

$$\frac{dS_t}{S_t} = (r - \tilde{\lambda}\bar{k})dt + \sigma dW_t + [Y_t - 1] dN_t,$$

where  $r$  is the constant risk-free interest rate,  $W_t$  is the standard  $\mathbb{Q}$ -Brownian motion,  $\bar{k} = \mathbb{E}[Y_t - 1]$ ,  $\ln Y_t \sim N(\alpha_J, \sigma_J^2)$ , and  $N_t$  follows the Poisson process with constant intensity  $\tilde{\lambda}$ . The first four moments of the log-return of  $S_t$ ,  $X_t = \ln(S_t/S_0)$  can be computed analytically (Ballotta and Kyriakou, 2015) as

$$\begin{aligned} \text{Mean} &= [r - \frac{\sigma^2}{2} - \tilde{\lambda}(e^{\alpha_J + \sigma_J^2/2} - 1) + \tilde{\lambda}\alpha_J]t \\ \text{Variance} &= (\sigma^2 + \tilde{\lambda}\alpha_J^2 + \tilde{\lambda}\sigma_J^2)t \\ \text{Skewness} &= \frac{\tilde{\lambda}(\alpha_J^3 + 3\alpha_J\sigma_J^2)}{\sqrt{t}(\sigma^2 + \tilde{\lambda}\alpha_J^2 + \tilde{\lambda}\sigma_J^2)^{3/2}} \\ \text{Kurtosis} &= \frac{\tilde{\lambda}(\alpha_J^4 + 6\alpha_J^2\sigma_J^2 + 3\sigma_J^4)}{t(\sigma^2 + \tilde{\lambda}\alpha_J^2 + \tilde{\lambda}\sigma_J^2)^2} + 3. \end{aligned} \tag{A.27}$$

The Johnson curve transformation (Johnson, 1949) transforms a standard normal variable into an arbitrary random variable via matching the first four moments. The nodes are set to be

$$X_i^n = \varepsilon g^{-1}\left(\frac{z_i - \gamma}{\delta}\right) + \nu, \tag{A.28}$$

where the parameters  $\gamma, \delta, \nu$  and  $\varepsilon$  can be determined by the algorithm proposed in Hill and Holder (1976),  $z_i$  are the discrete values of the standard normal distribution and the function  $g^{-1}(u)$  is defined by

$$g^{-1}(u) = \begin{cases} e^u & \text{for the lognormal family,} \\ \frac{e^u - e^{-u}}{2} & \text{for the unbounded family,} \\ \frac{1}{1 + e^{-u}} & \text{for the bounded family,} \\ u & \text{for the normal family.} \end{cases} \tag{A.29}$$

The  $m$  possible log-returns  $X_i^n$ ,  $i = 1, 2, \dots, m$ , are selected to match the first four moments of  $X_{t_n}$  by the Johnson curve transformation. The key in sampling  $X_i^n$  is to select  $\{z_i\}$  from the standard normal distribution. The corresponding underlying asset prices  $S_i^n$  on the wilow tree can then be calculated as  $S_i^n = S_0 \exp(X_i^n)$ .

The transition probability  $p_{ij}^n$  from  $X_i^n$  to  $X_j^{n+1}$  can be estimated by (Xu and Yin, 2014)

$$p_{ij}^n = P(A < X_j^{n+1} < B | X_i^n) = \int_{C_j^{n+1}} \sum_{l=0}^{C_{j+1}^{n+1}} \frac{e^{-\tilde{\lambda}\Delta t} (\tilde{\lambda}\Delta t)^l}{l!} \frac{1}{\sqrt{2\pi}\sigma_l} \exp\left[-\frac{(x - \mu_l)^2}{2\sigma_l^2}\right] dx, \tag{A.30}$$

where  $C_j^{n+1} = (X_{j-1}^{n+1} + X_j^{n+1})/2$ ,  $C_{j+1}^{n+1} = (X_{j+1}^{n+1} + X_j^{n+1})/2$ ,  $C_1^{n+1} = -\infty$ ,  $C_{m+1}^{n+1} = +\infty$ ,  $\mu_l = X_i^n + (r - \tilde{\lambda}\bar{k} - \sigma^2/2)\Delta t + l\alpha_J$  and  $\sigma_l^2 = \sigma^2\Delta t + l\sigma_J^2$ .

### A.3 Heston's stochastic volatility model

Assume the underlying asset price  $S_t$  follows a Heston stochastic volatility model (Heston, 1993)

$$\begin{cases} dS_t = \mu S_t dt + \sqrt{v_t} S_t dW_t^1, \\ dv_t = \eta(\varpi - v_t)dt + \sigma_v \sqrt{v_t} dW_t^2, \end{cases} \quad (\text{A.31})$$

where  $r$  is the risk-free interest rate,  $\varpi$  is the long-term mean of variance,  $\eta$  is a mean-reverting speed parameter of the variance,  $\sigma_v$  is the so-called volatility of volatility. The two Wiener processes  $dW_t^1$  and  $dW_t^2$  are assumed to be correlated with a constant correlation coefficient  $\rho$ , that is  $\mathbb{E}^\mathbb{Q}[dW_t^1 dW_t^2] = \rho dt$ . To ensure the variance is always positive, the Feller condition must be satisfied, i.e.,  $2\eta\varpi \geq \sigma_v^2$ .

When the Feller condition is satisfied, the details of a two-dimensional willow tree construction for  $S_t$  and  $v_t$  can be referred in Ma et al. (2020b). In summary, the willow tree consists of the tree nodes, that are pairs of possible values of  $S_t$  and  $v_t$ ,  $(S_i^n, v_{i_1}^n)$ , at time  $t_n = n \cdot \Delta t$ , where  $i = 1, 2, \dots, m$  and  $i_1 = \lfloor i/m_x \rfloor + 1$ ,  $m_x$  is the number of possible values of  $S_t$  at  $t_n$  given a possible value of  $v_t$  and  $\lfloor a \rfloor$  returns the largest integer less than  $a$ , and the transition probability between  $(S_i^n, v_{i_1}^n)$  at  $t_n$  and  $(S_j^{n+1}, v_{j_1}^{n+1})$  defined as  $p_{ij}^n$ .

When the Feller condition is violated, the variance in the Heston model could be negative in our willow tree or Monte Carlo simulation. In this paper, we adopt the adaptation in Cozma and Reisinger (2020) to set the negative variance to be zero. For example, at  $t_n \equiv n\Delta t$ , one of the variance on the willow tree,  $v_{i_1}^n$ , is negative, i.e.,  $v_{i_1}^n < 0$ . We first set it to be zero, i.e.,  $v_{i_1}^n = 0$ . Then, the transition probability from tree node at  $t_n$ ,  $(S_i^n, v_{i_1}^n)$ , to the tree nodes,  $(S_j^{n+1}, v_{j_1}^{n+1})$ , at  $t_{n+1}$ ,  $p_{ij}^n$ , can be determined as

$$p_{ij}^n = \begin{cases} 1 & \text{if } \ln S_i^n + r\Delta t \in \left[ \frac{\ln S_j^{n+1} + \ln S_{j-1}^{n+1}}{2}, \frac{\ln S_j^{n+1} + \ln S_{j+1}^{n+1}}{2} \right] \\ & \text{and } \eta\varpi\Delta t \in \left[ \frac{v_{j_1}^{n+1} + v_{j_1-1}^{n+1}}{2}, \frac{v_{j_1}^{n+1} + v_{j_1+1}^{n+1}}{2} \right] \\ 0 & \text{otherwise} \end{cases},$$

given  $v_{i_1}^n = 0$  for  $j = 1, 2, \dots, m$  and  $j_1 = \lfloor j/m_x \rfloor + 1$ .

## B Proof of Proposition 1

**Proof.** Equation (3.6) can be written as

$$\begin{aligned} \exp\left(-\frac{\theta}{1-R}(n+1)\Delta t\right) &= \mathbb{E}\left[\exp\left(-\sum_{k=0}^n \int_{t_k}^{t_{k+1}} \alpha(t, V)dt\right)\right] \\ &\approx \mathbb{E}\left[\exp\left(-\sum_{k=0}^n \alpha^k \Delta t\right)\right]. \end{aligned} \quad (\text{B.32})$$

where  $\alpha^k \triangleq \alpha(t_k, V^k)$ . Dividing both sides of the equation (B.32) by  $\exp(-\frac{\theta}{1-R}(n+1)\Delta t)$ , we have

$$\begin{aligned} 1 &= \mathbb{E} \left[ \exp \left( - \sum_{k=0}^n \alpha^k \Delta t + \frac{\theta}{1-R}(n+1)\Delta t \right) \right] \\ &= \mathbb{E} \left[ \prod_{k=0}^n \exp \left( -\alpha^k \Delta t + \frac{\theta}{1-R}\Delta t \right) \right] \\ &= \mathbb{E} \left[ \prod_{k=0}^n \eta^k \right], \end{aligned} \quad (\text{B.33})$$

where  $\eta^n \triangleq \exp \left( -\alpha^n \Delta t + \frac{\theta}{1-R}\Delta t \right)$ .

For  $n = 0$ , according to (B.33), we have

$$\mathbb{E}[\eta^0] = \exp \left( -\alpha^0 \Delta t + \frac{\theta}{1-R}\Delta t \right) = 1,$$

i.e.,  $\alpha^0 = \exp(a^0 + bV^0) = \frac{\theta}{1-R}$ .

For  $n = 1$ , we have

$$\mathbb{E}[\eta^0 \eta^1] = \mathbb{E}[\eta^1] = \sum_{i=1}^m q_i \eta_i^1 = 1.$$

Let  $w_i^1 = q_i$ , and we obtain equation (3.8).

For  $n > 1$ , on the one hand,  $\sum_{i=1}^m w_i^n \eta_i^n$  can be written as, based on the definition of  $w_i^n$

$$\begin{aligned} \sum_{i=1}^m w_i^n \eta_i^n &= \sum_{i=1}^m \eta_i^n \left[ \sum_{j_{n-1}=1}^m p_{j_{n-1}i}^{n-1} \eta_{j_{n-1}}^{n-1} w_{j_{n-1}}^{n-1} \right] \\ &= \sum_{i=1}^m \eta_i^n \left[ \sum_{j_{n-1}=1}^m p_{j_{n-1}i}^{n-1} \eta_{j_{n-1}}^{n-1} \left[ \sum_{j_{n-2}=1}^m p_{j_{n-2}j_{n-1}}^{n-2} \eta_{j_{n-2}}^{n-2} w_{j_{n-2}}^{n-2} \right] \right] \\ &= \sum_{i=1}^m \eta_i^n \left[ \sum_{j_{n-1}=1}^m p_{j_{n-1}i}^{n-1} \eta_{j_{n-1}}^{n-1} \cdots \left[ \sum_{j_1=1}^m p_{j_1j_2}^1 \eta_{j_1}^1 w_{j_1}^1 \right] \right]. \end{aligned} \quad (\text{B.34})$$

On the other hand, due to  $\eta^0 = 1$ ,  $\mathbb{E}[\prod_{k=0}^n \eta^k]$  can be expressed discretely in the willow tree framework as the sum of  $\prod_{k=1}^n \eta_{j_k}^k$  with probability  $q_{j_1} \prod_{k=1}^{n-1} p_{j_k j_{k+1}}^k$  for each  $j_1, j_2, \dots, j_n = 1, 2, \dots, m$ , i.e.,

$$\begin{aligned} \mathbb{E}[\prod_{k=1}^n \eta^k] &= \sum_{j_n=1}^m \cdots \sum_{j_1=1}^m \left( \prod_{k=1}^n \eta_{j_k}^k \right) (q_{j_1} \prod_{k=1}^{n-1} p_{j_k j_{k+1}}^k) \\ &= \sum_{j_n=1}^m \eta_{j_n}^n \left[ \sum_{j_{n-1}=1}^m p_{j_{n-1}j_n}^{n-1} \eta_{j_{n-1}}^{n-1} \cdots \left[ \sum_{j_1=1}^m p_{j_1j_2}^1 \eta_{j_1}^1 q_{j_1} \right] \right]. \end{aligned} \quad (\text{B.35})$$

Then, according to (B.34), (B.35) and (B.33), we have

$$\sum_{i=1}^m w_i^n \eta_i^n = \mathbb{E}[\prod_{k=0}^n \eta^k] = 1.$$

Thus, we have proved equation (3.8).

□

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